

# Frequentist network meta-analysis using the R package netmeta

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# Outline

Starting point: Graph-theoretical methods for network meta-analysis

Statistical model

Multi-arm studies

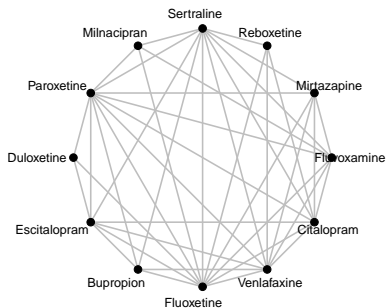
Drawing the network

Ranking treatments

Inconsistency diagnostics

Summary

# Graph-theoretical methods for network meta-analysis



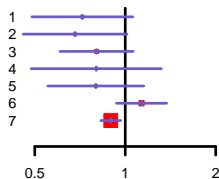
- ▶ Networks are **graphs**
  - ▶ **Nodes** are treatments
  - ▶ **Edges** are comparisons between treatments, based on studies
- ▶ '**Variiances combine like electrical resistances**' (Bailey, 2007)
- ▶ It is possible to apply methods from electrical network theory to network meta-analysis (Rücker, 2012)

# Variances combine like electrical resistances



- **Connection in series** Variances in a chain of  $n - 1$  independent comparisons of successive treatments  $A, B, C, \dots$  add:

$$V_{A-E} = V_{A-B} + V_{B-C} + V_{C-D} + V_{D-E}$$



- **Parallel connection** For a pairwise meta-analysis with parallel comparisons, inverse variances add:

$$\frac{1}{V(\bar{x})} = \sum_k \frac{1}{V_k}$$

# Terminology in meta-analysis and electrical networks

## Meta-analytic network

## Electrical network

Treatments $i = 1, \dots, n$	$\iff$	Nodes $i = 1, \dots, n$
Existing comparisons $e = 1, \dots, m$	$\iff$	Edges $e = 1, \dots, m$
Variance $V_e$	$\iff$	Resistance $R_e$
Inverse variance weight $w_e = 1/V_e$	$\iff$	Conductance $1/R_e$
Outcome of treatment $i$	$\iff$	Potential at node $i$
Treatment effect $i - j$	$\iff$	Voltage at edge $i - j$
Weighted treatment effect $i - j$	$\iff$	Current flow at edge $i - j$

- ▶ Ohm's law relates treatment effects and weights
- ▶ Kirchhoff's current law says how to combine the observed effects
- ▶ Kirchhoff's potential law guarantees **consistency** of the estimated treatment effects over closed circuits
  - ▶ Consistency means that the difference between two treatments is always the same, whatever (direct or indirect) path is chosen

# Statistical model

## Model

$$\hat{\boldsymbol{\theta}} = \mathbf{X}\boldsymbol{\theta}^{treat} + \boldsymbol{\epsilon}, \quad \boldsymbol{\epsilon} \sim N(\mathbf{0}, \boldsymbol{\Sigma}),$$

## where

- ▶  $\hat{\boldsymbol{\theta}}$  is a vector of  $m$  observed pairwise comparisons with known standard errors  $\mathbf{s} = (s_1, s_2, \dots, s_m)$
- ▶  $\mathbf{X}$  is the  $m \times n$  design matrix defining the network structure
- ▶  $\boldsymbol{\theta}^{treat}$  a vector of length  $n$  (number of treatments)
- ▶  $\boldsymbol{\Sigma}$  is a diagonal matrix whose  $i^{th}$  entry is  $s_i^2$ .

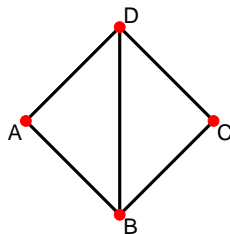
## Note:

- ▶ If there are  $K$  two-arm trials,  $\hat{\boldsymbol{\theta}}$  has length  $K$
- ▶ If there are also multi-arm trials,  $\hat{\boldsymbol{\theta}}$  has length  $m \geq K$  with  $m$  denoting the total number of pairwise comparisons

## Example network with $n = 4$ arms

Example network with  $n = 4$  arms

- ▶  $\theta^{treat} = (\theta_A, \theta_B, \theta_C, \theta_D)^T$
- ▶  $K = 5$  studies each providing a single pairwise treatment comparison
- ▶  $m = 5$  pairwise treatment comparisons
- ▶ Model:



$$\begin{pmatrix} \hat{\theta}_1^{AB} \\ \hat{\theta}_2^{BC} \\ \hat{\theta}_3^{CD} \\ \hat{\theta}_4^{AD} \\ \hat{\theta}_5^{BD} \end{pmatrix} = \begin{pmatrix} 1 & -1 & 0 & 0 \\ 0 & 1 & -1 & 0 \\ 0 & 0 & 1 & -1 \\ 1 & 0 & 0 & -1 \\ 0 & 1 & 0 & -1 \end{pmatrix} \begin{pmatrix} \theta_A \\ \theta_B \\ \theta_C \\ \theta_D \end{pmatrix} + \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \\ \epsilon_5 \end{pmatrix} \\
 = \mathbf{X}\theta^{treat} + \epsilon$$

## Estimation under the fixed effect model

- ▶  $\mathbf{W} = \text{diag}(1/s_1^2, \dots, 1/s_m^2)$  diagonal matrix (dimension  $m \times m$ ) of inverse variance weights
- ▶ Network estimates  $\hat{\theta}^{nma}$  estimated by

$$\hat{\theta}^{nma} = \mathbf{H}\hat{\theta}$$

where  $\mathbf{H} = \mathbf{X}(\mathbf{X}^T\mathbf{W}\mathbf{X})^+\mathbf{X}^T\mathbf{W}$  is known as the *hat matrix* in regression.

- ▶ **Interpretation: The network estimates are weighted sums of the observed estimates with weights coming from the rows of  $\mathbf{H}$ .**
- ▶ Standard errors calculated from the variance-covariance matrix

$$\widehat{\text{Cov}}(\hat{\theta}^{nma}) = \mathbf{X}(\mathbf{X}^T\mathbf{W}\mathbf{X})^+\mathbf{X}^T$$

- ▶ Heterogeneity/inconsistency measured by generalised  $Q_{total}$  statistic

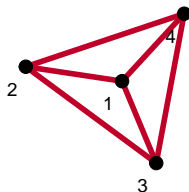
$$Q_{total} = (\hat{\theta} - \hat{\theta}^{nma})^T \mathbf{W} (\hat{\theta} - \hat{\theta}^{nma})$$

(Jackson et al., 2012; Rücker, 2012; Krahn et al., 2013)



# Multi-arm studies: Need to account for correlation

- ▶ A study with  $k$  arms contributes  $\binom{k}{2}$  pairwise comparisons
- ▶ **Note: These are correlated, as there are only  $k$  treatments**
  - ▶  $k - 1$  independent comparisons
  - ▶  $k - 1$  degrees of freedom ( $df$ )
- ▶ Example  $k = 4$ :  $df = 3$



# Adjustment for correlation within multi-arm studies

## Standard approach: Reduce dimension

(Lu et al., 2011; Higgins et al., 2012; White et al., 2012; König et al., 2013)

- ▶ Based on standard regression methodology
- ▶ For each multi-arm study, choose a study-specific reference treatment
- ▶ Consider only comparisons to the reference treatment ('basic parameters')

## Alternative approach: Reduce weights

(Rücker, 2012; Rücker and Schwarzer, 2014)

- ▶ Based on electrical network methodology
- ▶ For each multi-arm study, reduce all 'conductances' (weights) by specific factors that must be calculated
- ▶ Implemented in the R package **netmeta** (Rücker et al., 2016)

# Comparison of the approaches

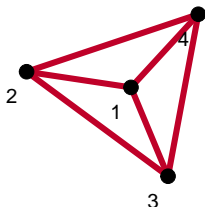
## Standard approach

- ▶ Natural for statisticians with a background in regression analysis

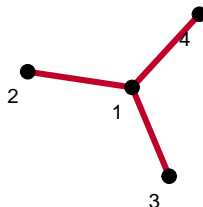
## Alternative approach

- ▶ Natural for scientists coming from graph theory and its applications

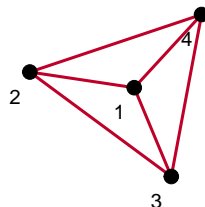
Given a four-arm study with six comparisons,



we may cut off three of six comparisons:



or reduce all weights by 1/2 (in average):



# Examples

## 1. Diabetes data

Network of 10 diabetes treatments including 26 studies, where the outcome was HbA1c (measured as mean change or mean post treatment value) (Senn et al., 2013)

## 2. Smoking cessation data

Network of four interventions for smoking cessation (binary outcome) (Higgins et al., 2012; Dias et al., 2013)

Both examples are part of R package **netmeta**

# How to use R package netmeta: Diabetes data

```
# Make R package netmeta available  
install.packages("netmeta")  
library(netmeta)
```

```
# Load diabetes data (Senn 2013), included in R package netmeta  
data(Senn2013)  
# Look at first 5 lines: data are in contrast-based format  
head(Senn2013, 5)
```

```
##      TE   seTE treat1 treat2      studlab  
## 1 -1.90 0.1414  metf  plac  DeFronzo1995  
## 2 -0.82 0.0992  metf  plac    Lewin2007  
## 3 -0.20 0.3579  metf  acar   Willms1999  
## 4 -1.34 0.1435  rosi  plac   Davidson2007  
## 5 -1.10 0.1141  rosi  plac Wolffenbuttel1999
```

```
# Network meta-analysis of diabetes data  
net1 <- netmeta(TE, seTE, treat1, treat2, studlab, data = Senn2013, sm = "MD",  
  comb.fixed=FALSE, comb.random=TRUE)
```

# Summary output of diabetes data

```

# Summarize results
summary(net1)

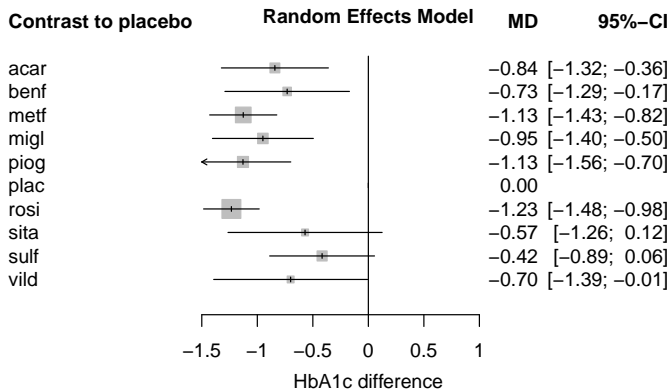
## Number of studies: k=26
## Number of treatments: n=10
## Number of pairwise comparisons: m=28
##
## Random effects model
##
## Treatment estimate (sm='MD'):
##      acar      benf      metf      migl      piog      plac      rosi      sita
## acar      . -0.1106  0.2850  0.1079  0.2873 -0.8418  0.3917 -0.2718
## benf  0.1106      .  0.3956  0.2186  0.3979 -0.7311  0.5023 -0.1611
## metf -0.2850 -0.3956      . -0.1770  0.0023 -1.1268  0.1067 -0.5568
## migl -0.1079 -0.2186  0.1770      .  0.1794 -0.9497  0.2837 -0.3797
*** Output truncated ***
##
## Quantifying heterogeneity/inconsistency:
## tau^2 = 0.1087; I^2 = 81.4%
##
## Test of heterogeneity/inconsistency:
##      Q d.f.  p-value
## 96.99  18 < 0.0001

```

# Forest plot of diabetes data

```
# Look at result
```

```
forest(net1, ref = "plac",
       pooled = "random", digits=2,
       smlab = "Random effects model",
       xlab = "HbA1c difference",
       leftlabs = "Contrast to placebo")
```



# Smoking cessation data

```
# Load diabetes data (Senn 2013)
```

```
data(smokingcessation)
```

```
# Look at first lines: data are in arm-based format
```

```
head(smokingcessation)
```

```
##      event1  n1 event2   n2 event3  n3 treat1 treat2 treat3
## 1         9 140     23  140     10  138      A      C      D
## 2        11  78     12   85     29 170      B      C      D
## 3        75 731    363  714     NA  NA      A      C
## 4         2 106      9  205     NA  NA      A      C
## 5        58 549    237 1561     NA  NA      A      C
## 6         0  33      9   48     NA  NA      A      C
```

```
# The first two trials are three-arm trials
```



# Smoking cessation data

```
# Transform data from arm-based format to contrast-based format
```

```
p2 <- pairwise(treat = list(treat1, treat2, treat3),
               event = list(event1, event2, event3),
               n = list(n1, n2, n3),
               data = smokingcessation, sm = "OR")
```

```
head(p2, 9)
```

```
##           TE      seTE studlab treat1 treat2 event1  n1 event2  n2
## 1 -1.051293027 0.4132432      1     A     C      9 140     23 140
## 2 -0.128527575 0.4759803      1     A     D      9 140     10 138
## 3  0.922765452 0.3997972      1     C     D     23 140     10 138
## 4 -0.001244555 0.4504070      2     B     C     11  78     12  85
## 5 -0.225333286 0.3839393      2     B     D     11  78     29 170
## 6 -0.224088731 0.3722995      2     C     D     12  85     29 170
## 7 -2.202289286 0.1430439      3     A     C     75 731    363 714
## 8 -0.870353637 0.7910933      4     A     C      2 106      9 205
## 9 -0.415648522 0.1557329      5     A     C     58 549    237 1561
```

```
# Note the two three-arm studies 1 and 2, now each filling three data lines
```

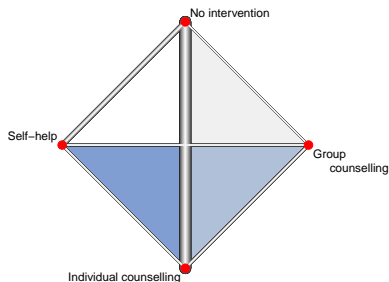
# Smoking cessation data

```
net2 <- netmeta(TE, seTE, treat1, treat2, studlab, data = p2,  
               comb.fixed = FALSE, comb.random = TRUE)  
summary(net2)
```

```
## Number of studies: k=24  
## Number of treatments: n=4  
## Number of pairwise comparisons: m=28  
##  
## Random effects model  
##  
## Treatment estimate (sm='OR'):  
##           A           B           C           D  
## A           . 0.6595 0.4803 0.4056  
## B 1.5162           . 0.7282 0.6150  
## C 2.0822 1.3732           . 0.8446  
## D 2.4653 1.6259 1.1840           .  
*** (Output truncated) ***  
  
## Quantifying heterogeneity/inconsistency:  
## tau^2 = 0.5989; I^2 = 88.6%  
## Test of heterogeneity/inconsistency:  
##           Q d.f. p.value  
## 202.62 23 < 0.0001
```

# Smoking cessation data

```
# Define treatment names  
tname <- c("No intervention", "Self-help", "Individual counselling", "Group  
counselling")  
# Produce network graph  
# Transparent coloured areas correspond to three-arm studies  
netgraph(net2, points=TRUE, cex.points=3, cex=1.25, labels=tname)
```

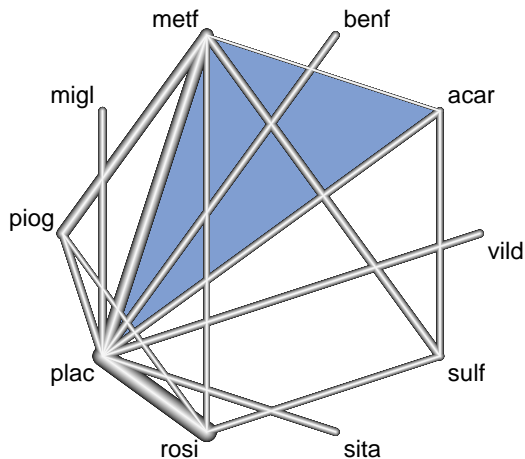


## Drawing the network with **netmeta**

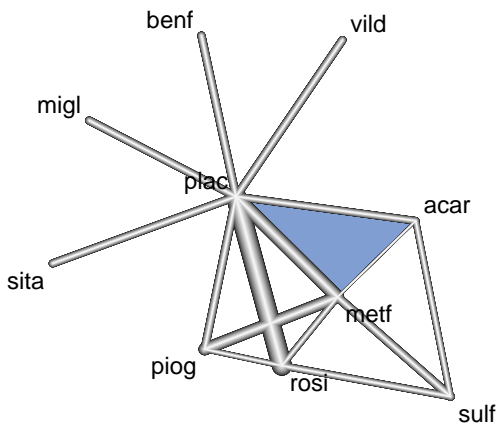
For network visualisation, use function **netgraph**

- ▶ Iteration method implemented in **netmeta**: **Stress algorithm** (Kamada and Kawai, 1989; Hu, 2012, related to multi-dimensional scaling)
- ▶ Various starting (also random) layouts available
- ▶ Iteration steps visible/printable, if desired
- ▶ Variable choice of scale, node size, line width, colours, highlighting
- ▶ Coloured polygons may represent multiarm studies (where transparent colours are available)

# Drawing the network with netmeta: Diabetes data



# Drawing the network with netmeta: Diabetes data



# Ranking treatments

- ▶ **Bayesian framework:**

Derive ranking probabilities for each treatment from the posterior distributions

- ▶ Treatments may be ranked by the surface under the cumulative ranking curve (SUCRA) (Salanti et al., 2011)

- ▶ **Frequentist framework:**

We introduced a quantity, called P-score, as an analogue to SUCRA (Rücker and Schwarzer, 2015)

- ▶ Example: Diabetes data





## Ranking treatments using P-scores: Diabetes data

- ▶ **P-scores** allow ranking the treatments on a continuous 0-1 scale
- ▶ Based on frequentist point estimates and standard errors
- ▶ Frequentist analogue to SUCRA (Rücker and Schwarzer, 2015)

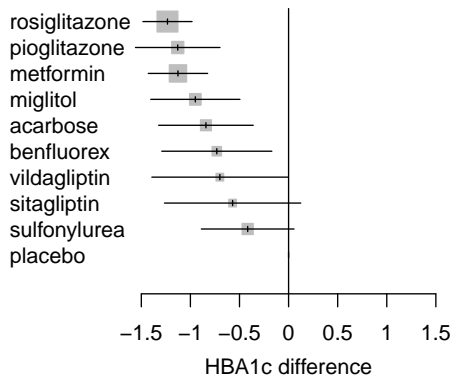
```
# Rank treatments
# Small values are "good" here (this is the default), otherwise "bad"
netrank(net1, small.values = "good")
```

```
##      P-score
## rosi  0.8934
## metf  0.7818
## piog  0.7746
## migl  0.6137
## acar  0.5203
## benf  0.4358
## vild  0.4232
## sita  0.3331
## sulf  0.2103
## plac  0.0139
```

# Ranking treatments using P-scores: Diabetes data

Compare forest plot, point estimates, SUCRA values and P-scores

**Treatment**      **REM (frequentist analysis)**



	Frequentist	SUCRA	P-score
rosi	-1.23	0.890	0.893
metf	-1.13	0.780	0.782
piog	-1.13	0.773	0.775
migl	-0.95	0.620	0.614
acar	-0.84	0.520	0.520
benf	-0.73	0.439	0.436
vild	-0.70	0.413	0.423
sita	-0.57	0.334	0.333
sulf	-0.42	0.213	0.210
plac	0	0.018	0.014

# Inconsistency diagnostics

## Designs in network meta-analysis

- ▶ A **design** is each combination of treatments within a study in a network meta-analysis
  - ▶ Example: For three treatments  $A, B, C$ , the possible designs are  $A : B, A : C, B : C, A : B : C$
  - ▶ For  $n$  treatments the maximum number of designs is  $2^n - n - 1$
  - ▶ Not all these need be present in a given network meta-analysis
  - ▶ In a pairwise meta-analysis, all trials have the same design  $A : B$
- ▶ **Clinical context**
  - ▶ Example: Studies with design  $A : C$  might differ to studies with design  $A : B$  or  $A : B : C$  in that they include patients who cannot be randomised to  $B$
  - ▶ Heterogeneity between designs is plausible

# Decomposition of the heterogeneity statistic

Total Q statistic

$$Q_{total} = (\hat{\theta} - \hat{\theta}^{nma})^T \mathbf{W} (\hat{\theta} - \hat{\theta}^{nma})$$

Krahn et al. (2013):

- ▶ Q can be decomposed into
  - ▶ a part coming from **within designs** (heterogeneity between studies of the same design)
  - ▶ a part coming from **between designs** (inconsistency between studies of different designs)
- ▶ Q can be decomposed into parts coming from each design
- ▶ Q can be decomposed into parts coming from each study

# Decomposition of Q: Diabetes data

```
# Decompose total Q statistics into parts from designs
decomp.design(net1)

## Q statistics to assess homogeneity / consistency
##
##           Q df  p.value
## Whole network  96.99 18 < 0.0001
## Within designs  74.46 11 < 0.0001
## Between designs 22.53  7  0.0021
##
## Design-specific decomposition of within-designs Q statistic
##
##      Design      Q df  p.value
##      acar:plac  0.00  0      --
##      acar:sulf  0.00  0      --
##      benf:plac  4.38  1  0.0363
##      metf:piog  0.00  0      --
##      metf:plac 42.16  2 < 0.0001
##      metf:rosi  0.19  1  0.6655
##      metf:sulf  0.00  0      --
*** (Output truncated) ***
##      acar:metf:plac  0.00  0      --
```

# Decomposition of Q: Diabetes data

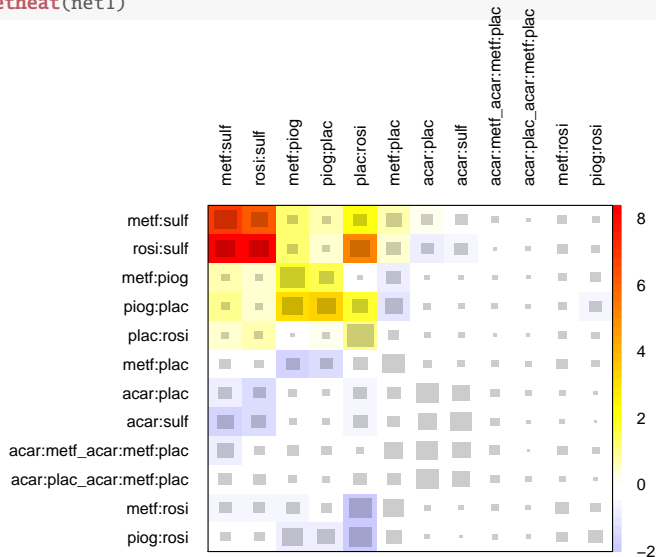
```
# Decompose total Q statistics into parts from designs
decomp.design(net1)

## Between-designs Q statistic after detaching of single designs
##
## Detached design      Q df  p.value
##      acar:plac  22.44  6  0.001
##      acar:sulf  22.52  6  0.001
##      metf:piog  17.13  6  0.0088
##      metf:plac  22.07  6  0.0012
##      metf:rosi  22.52  6  0.001
##      metf:sulf   7.51  6  0.276   ***
##      piog:plac  17.25  6  0.0084
##      piog:rosi  22.48  6  0.001
##      plac:rosi  16.29  6  0.0123
##      rosi:sulf   6.77  6  0.3425   ***
##      acar:metf:plac 22.38  5  0.0004
```

Explanation: Detaching a design means relaxing the consistency assumption for this design. If Q decreases markedly after detaching a design (\*\*\*) added for the purpose of this talk), we conclude that this design contributed to between-design inconsistency. If Q does not decrease markedly, the design is not thought to contribute to between-design inconsistency.

# Net heat plot (Krahn et al., 2013): Diabetes data

`netheat`(net1)



## Net heat plot (Krahn et al., 2013)

- ▶ Areas of grey squares ■: indicate the contribution from the treatment comparison in the column to the treatment comparison in the row
- ▶ Colours on the diagonal represent the inconsistency contribution of the corresponding design (**red** means large)
- ▶ Colours on the off-diagonal associated with the change in inconsistency between direct and indirect evidence in a network estimate in the row after relaxing the consistency assumption for the effect of one design in the column
  - ▶ **Blue** indicates that the evidence of the design in the column supports the evidence in the row
  - ▶ **Red** indicates that the evidence of the design in the column contrasts to the evidence in the row
- ▶ Largest inconsistency contribution by the `metf:sulf` and `rosi:sulf` designs (red squares in top left corner)



# Summary

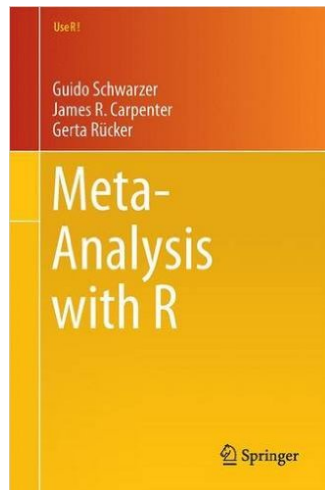
R package **netmeta** provides

- ▶ flexible data entry (**pairwise**)
- ▶ fixed / random effects model (**netmeta**)
- ▶ appropriate incorporation of multi-arm trials
- ▶ forest plots (**forest**)
- ▶ network graphs (**netgraph**)
- ▶ ranking of treatments (**netrank**)
- ▶ inconsistency diagnostics (**decomp.design**, **netheat**)

Currently not available:

Meta-regression

See book Schwarzer et al. (2015)



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## Appendix: A proof that SUCRA and P-score are the same

We assume the true probabilities as known. If  $R(i) = k$  means that treatment  $i$  has rank  $k$ , we have

$$P_{ij} = \sum_{k=1}^{n-1} \sum_{l=k+1}^n P(R(i) = k \wedge R(j) = l)$$

and

$$(n-1)SUCRA(i) = \sum_{r=1}^{n-1} F(i, r) = \sum_{r=1}^{n-1} \sum_{k=1}^r P(i, k) = \sum_{k=1}^{n-1} \sum_{r=k}^{n-1} P(i, k) = \sum_{k=1}^{n-1} (n-k)P(i, k)$$

It follows

$$\begin{aligned} \sum_{j=1}^n P_{ij} &= \sum_{j=1}^n \sum_{k=1}^{n-1} \sum_{l=k+1}^n P(R(i) = k \wedge R(j) = l) = \sum_{k=1}^{n-1} \sum_{l=k+1}^n \sum_{j=1}^n P(R(i) = k \wedge R(j) = l) \\ &= \sum_{k=1}^{n-1} \sum_{l=k+1}^n P(i, k) = \sum_{k=1}^{n-1} (n-k)P(i, k) = (n-1)SUCRA(i) \end{aligned}$$

and thus

$$\bar{P}_i = \frac{1}{n-1} \sum_{j=1}^n P_{ij} = SUCRA(i)$$

which is what we wanted to prove. Note: For  $n > 2$ , neither ranking probabilities  $P(i, k)$  nor probabilities  $P_{ij}$  can be uniquely determined from  $\bar{P}_i$  or  $SUCRA(i)$ .